

Who Benefits from R&D Tax Policy? *

¿Quién se beneficia de los incentivos fiscales a la inversión en I+D? *

M. Beatriz Corchuelo **
Ester Martínez-Ros ***

Sumario: 1. Introduction. 2. Literature Review. 3. Spanish Data. 4. Methods. 5. Empirical Results. 6. Conclusions. 7. References.

Recepción del original: 30/09/2009

Aceptación del original: 15/11/2010

ABSTRACT: This paper explores which types of firms benefit more from the R&D fiscal incentives and the effect of this type of instrument on the performance of R&D activities. Spain is considered the most generous countries in the OECD in fiscal treatment of R&D, yet our data reveal that tax incentives are little known and, especially, seldom used by firms. Restricting our empirical analysis to those firms that do report knowing about such incentives, we investigate the average effect of tax incentives on innovation, using both nonparametric methods (matching estimators) and parametric methods (Heckman's two-step selection model). First, we find that large firms, especially those that implement innovations, are more likely to use the tax incentives, while small and medium enterprises encounter some obstacles to using them. Secondly, the average effect of the policy is positive, but significant only in large firms. Our main conclusion is that tax incentives increase innovative activities by large and high-tech sector firms, but may be used only randomly by small medium enterprises.

Key words: : R&D fiscal incentives, matching methods

JEL Classification: O31, H25, H32.

RESUMEN: Este trabajo analiza qué tipo de empresas se benefician más de los incentivos fiscales a la inversión en I+D y el efecto que tienen sobre la inversión en I+D. España se considera el país más generoso en el tratamiento fiscal de este tipo de inversión si bien los

* We are grateful for comments from Isabel Busom (Universidad Autónoma of Barcelona), José María Labeaga (IEF, Spain), Pierre Mohnen (UNU-MERIT University, Maastricht) and Joaquín Artés (Universidad Complutense, Madrid) and from seminar participants at the Applied Economics Meeting in Logroño (Spain), the Symposia of Economics Analysis in Granada (Spain) and the Public Economics Meeting in Salamanca (Spain). We acknowledge financial support from project n° 2007/4162/001 provided by the Spanish Fiscal Studies Institute, project n° 03963 provided by Education and Science Spanish Ministry and project n° 0059 provided by Comunidad de Madrid.

** Universidad de Extremadura. Avda. de Elvas, s/n. 06071 Badajoz. Tfo: 924289300. Fax: 924272509. e-mail: bcorchue@unex.es

*** Universidad Carlos III de Madrid. C/. Madrid, 126-128. 28903 Getafe. Tfo. 91 624 9350. Fax: 91 624 9607. e-mail: emros@emp.uc3m.es

datos muestran que los incentivos fiscales se conocen poco y, sobre todo, que se aplican escasamente. Restringiendo nuestro análisis a la submuestra de empresas que declaran conocer los incentivos fiscales, se investiga el efecto promedio de los incentivos fiscales sobre la inversión en I+D utilizando métodos econométricos paramétricos (*modelo de selección en dos etapas de Heckman*) y no paramétricos (*estimadores matching*). Obtenemos que, en primer lugar, las empresas grandes (especialmente las innovadoras) son las que tienen más probabilidad de aplicar los incentivos fiscales, mientras que las pequeñas y medianas empresas (pymes) encuentran más dificultades en su aplicación. En segundo lugar, el efecto promedio de la política es positivo, si bien solo estadísticamente significativo en las grandes empresas. La principal conclusión del estudio es que los incentivos fiscales se aplican aleatoriamente en las pymes pero no en las grandes empresas y aquellas que operan en sectores de alta-media intensidad tecnológica que valoran sus beneficios con un efecto positivo en sus actividades de innovación.

Palabras clave: Incentivos fiscales a la inversión en I+D, estimadores matching.

Clasificación JEL: O31, H25, H32.

1. Introduction

Evidence suggests that private firms spend less on R&D than is socially optimal because the presence of externalities creates a gap between private and public profitability (Arrow 1962; Nelson 1959). In most OECD countries governments try to solve this problem through various measures, such as subsidies or fiscal incentives for R&D, in order to stimulate innovation. The effectiveness of these instruments depends on each country's fiscal system and the purposes of its programs (European Commission 2003).

Spain is considered the most generous country in the OECD in tax treatment of R&D (Warda 2001 and 2002). Fiscal benefits act either on the tax base (free depreciation) or on the corporate income tax liability (tax credits for R&D and technological innovation) and apply to all firms that carry out R&D activities, independently of the success or failure of the project¹. Given these tax incentives, the price of R&D is minor (Marra 2004; Corchuelo 2006), so these instruments should in theory stimulate innovation activities and correct for suboptimal R&D funding in firms.

However, empirical evidence reveals that these generous tax incentives are rarely used by Spanish firms (Corchuelo and Martínez-Ros, 2008). Accordingly, the main goal of this paper is to evaluate who benefits from this policy and the effectiveness of tax incentives for R&D in Spain. The success of this tax policy is based on the government's ability to design it and on the use of firms' expected profit. Therefore, it is very important to understand the design of the tax benefits and the characteristics of the firms that use them. This means that the variable that explains the financial-economic profitability of the incentives (generally measured by the marginal effective tax credit, B-index or cost of R&D capital) is endogenous. We assume also that there may be a selection problem—that is, the design of the tax incentives may make some kinds of firms more likely to use them than others.

¹ Fiscal incentives for R&D in Spain are governed by articles 11.2, 16.4, 35, 36 and 40.3 of the Revised Text of the Corporate Income Tax Law (RD-L 4/2004).

We find that firms with higher capacity for innovation, with a stable financial position, and that have also received R&D subsidies (especially SMEs) are more likely to take advantage of the tax benefits. However, SMEs encounter important obstacles to using them. In addition to this, when we empirically analyse the effect of this tax policy, we find that this type of incentives has a significant and important effect only on large firms and firms that belong to high-technological intensity sectors and it is not significant for SMEs which do not suffer from selection problems. Then, we conclude that tax incentives for R&D are randomly distributed among SMEs because those firms are not aware of their benefits. In contrast, large firms consider tax incentives as a possible instrument to capture the returns from R&D investment because it reduces its costs and the policy has a positive effect on their technological effort. Empirical studies that have analyzed the effectiveness of tax incentives for R&D have usually employed parametric techniques of econometric estimation. This paper uses a nonparametric approach (matching methods) to take into account the problems of selection and endogeneity. We also present a parametric approach (Heckman's two-step selection model) to compare our results with those from the standard approach.

The matching approach has been widely used in the evaluation of public policies, especially policies oriented to the labour market (Heckman *et ál.*, 1997; Heckman *et ál.*, 1999; Lechner, 2002; Dehejia and Wahba, 1999 and 2002; Hotz *et ál.*, 2006). In the evaluation of innovation policies, this method has been used, essentially, to analyse the effects of direct subsidies on R&D activities; little research has addressed the effects of tax incentives (Czarnitzki *et ál.* 2004; Heijs *et ál.* 2006). Also, only Marra (2004), Corchuelo (2006) and Romero and Sanz (2007) have studied the effects of tax incentives on innovation in Spain.

Our data come from the *Spanish Business Strategy Survey* at the firm level. Since 2001, this survey has included some questions related to firms' knowledge and use of tax incentives for R&D. Our analysis is focused on a single year, 2002, although we use variables covering the period 1998-2002.

Our results are presented in three parts. First, we isolate the factors that affect the probability that a firm will take advantage of tax incentives for R&D; second, we estimate the average effect of the tax policy on user and non-user firms, using a nonparametric approach (matching estimators); and third, we analyse the effects of tax benefits on R&D technological effort, using a parametric approach (Heckman's two-step selection model) in order to compare the results with those from the nonparametric method.

The rest of the paper is organized as follows: Section 2 reviews the literature evaluating innovation policies, Section 3 introduces the data sources, and Section 4 discusses the methods used for the empirical assessment of the effects. The empirical results are presented in Section 5, and Section 6 concludes.

2. Literature Review

There is abundant empirical literature related to the effects of financial public policies—either subsidies or tax incentives—on R&D spending. Studies of tax incentives aim mainly at measuring the additional private R&D spending and the cost-effectiveness of the incentives, that is, whether the additional R&D spending is greater than the government's tax revenue losses.

Most studies have used a micro-econometric framework because more and better micro-data are available, and most studies have focused on the American and Canadian economies (Hall and Van Reenen 2000). Nevertheless, there are also studies focused on other countries². Studies using a macro-econometric framework are scarce but not irrelevant, since macro-data allow capturing the indirect effects caused by the incentives, for example, the spillover effects of innovation between countries (Guellec and van Pottlesberghe 2003; Falk 2004), or other macro-economic factors that affect R&D decisions as the culture of innovation, the fiscal system, etc. (Bloom *et al.* 2002).

The methodology used in the literature on R&D fiscal incentives combines event studies comparing behaviour before and after a change in the policy (Collins 1983); questionnaire surveys and interviews attempting to determine how individual firms respond to a policy change (Mansfield 1986); and econometric estimation using two types of parametric methods: impact models (Berger 1993) in which a binary variable shows the impact of the tax incentive, and demand models (Hall 1993) that directly obtain the price-elasticity of R&D investment. Although results are mixed, the most recent studies confirm the effectiveness (Hall and van Reenen 2000)³ and cost-effectiveness (Finance Department of Canada 1998) of tax incentives for R&D. In Spain, there is little evidence of the impact of this tax policy, except for studies by Marra (2004), who concludes that tax incentives affect the firm cost structure and improve the private demand for R&D, and Corchuelo (2006), who finds that incentives increase the probability of innovation and R&D effort.

Recently, parametric approaches have been combined with nonparametric ones in the evaluation of public policies. This approach has been widely used to determine whether R&D direct subsidies complement or crowd out private spending on R&D. In these studies, the recent concern about endogeneity and selection problems has led researchers to identify patterns in the distribution of public support and to estimate counterfactual status (Herrera and Heijs 2006).

Using a matching approach, the studies of Czarnitzki (2001), Fier (2002), Czarnitzki and Fier (2002) and Almus and Czarnitzki (2003) in Germany; Duguet (2004) in France; Lööf and Hesmati (2005) in Sweden, and Czarnitzki and Ebersberger (2006) in Finland obtain that, on average, R&D subsidies do not fully crowd out private financing because innovation input increases with the re-

² Some examples: Lattimore (1997) in Australia, Asmussen and Berriot (1993) in France, van Pottlesberghe *et al.* (2003) in Belgium, and Parisi and Sembenelli (2003) in Italy.

³ In general, the median of the elasticity is -0.85, and the average is -0.81 (European Commission 2003).

ception of public subsidies. In Spain, Herrera and Heijs (2006) and González and Pazó (2008) reject either full or partial crowding out of private financing by R&D subsidies although the second one concludes that some firms (especially SMEs and firms with medium-low technological intensity) would have not developed R&D activities in the absence of R&D subsidies.

Other studies combine parametric and nonparametric techniques to evaluate public funding policy as Kaiser (2004) who uses Heckman's two-step selection model and nearest neighbour matching estimator, Görg and Strolb (2005) who work with a sample of Irish plants and using parametric (DID) and nonparametric estimations (with Caliper matching estimator), Aerts and Czarnitzki (2005) who combine matching estimators and instrumental variables and Duch et ál. (2007), in Spain, who use the matching and OLS approach to analyze the impact of various types of R&D subsidies received by Catalanian manufacturing and service sector firms. All of them reject the existence of a crowding-out effect as well.

Similar conclusions have been obtained in studies that have used only parametric methods of estimation to evaluate the relationship between R&D subsidies and firms' R&D activities as in Hussinger (2006) applied to the German economy; Lach (2002) in Israel; Hanel (2003) in Canada; Ali-Yrkkö (2005 a, b), in Finland; González et ál. (2005) and Gelabert *et ál.* (2009) in Spain.

To summarize, the majority of studies that have used only parametric or only nonparametric approaches or both, controlling for the endogeneity of R&D subsidies and sample selection problems, have rejected full crowding-out effects. The exceptions are Wallsten (2000) who finds full crowding out of private R&D spending in firms benefiting from the U.S. Small Business Innovation Research (SBIR) programme; Busom (2000), in Spain, who rejects full crowding out of private funds but finds partial crowding out for 30 percent of the firms; Toivanen and Niininem (2000), in Finland, who show that the positive effect of R&D subsidies disappears in firms with higher cash-flow; and Suetens (2002), in Belgium, who finds no significant effects of R&D subsidies.

In contrast, there are few papers that have analysed the effects of tax incentives on R&D using a nonparametric approach. In Canada, Czarnitzki *et ál.* (2004), using data for 1999, show that tax credits increase a firm's probability of developing new products in domestic and foreign markets; they infer that firms perform more R&D activities because of the tax incentives. In Spain, Heijs *et ál.* (2006) use the *propensity score matching* method to analyse the effect of tax incentives on Spanish manufacturing firms. They reject a crowding-out effect, although they conclude that fiscal incentives do not stimulate additional R&D investment. Lastly, Berube and Mohnen (2009) have analysed the effectiveness of R&D grants for Canadian plants already benefit from R&D tax credits concluding that these firms introduced more new products, more world-first product as well, and were more successful in commercializing their innovations than their counterfactual part that only benefited from R&D tax credit.

To sum up, the combination of parametric and nonparametric approach in the analysis of the effectiveness of tax incentives and their application to the Spanish economy where, as we previously regarded, it is still scarce, are the main contributions of this paper.

3. Spanish Data

We use data provided by the SEPI Foundation on behalf of the Spanish Ministry of Industry, Tourism and Trade⁴. This survey contains data from 1990 on manufacturing firms with more than 10 employees, which are questioned about their strategies in the short and long run, including their R&D decisions. Since 2001, this survey has included some questions about the knowledge and use of tax incentives for R&D that allow us to distinguish between firms that merely are aware of these incentives and firms that actually use them. Because of some problems with answers in 2001, our research refers to the year 2002, although we measure some variables for the period 1998 to 2002. Our final sample includes 1708 observations at the firm level.

Tables 1 and 2 analyse the knowledge and use of tax incentives by firms of various sizes⁵ and sectors⁶. Only a little more than half of the firms in the full sample are aware of the incentives; awareness is especially low among SMEs and in industries with medium-low technological intensity. And of those that know about the R&D tax credit, fewer than half use it. If we consider only firms that conducted R&D in 2002, the percentages increase, especially of SMEs and firms with medium-low technological intensity that report knowing about the tax incentives, but still, very few firms report using them. Finally, Table 3 analyses the average R&D technological effort (R&D spending over sales), which, in general, is higher in those firms that know about and, especially, that use the tax incentives. Whether there is a causal connection is something to confirm by empirical study.

TABLE 1.—*Firms That Know about and Use R&D Tax Incentives, by Size (2002)*

Full sample						
	Do not know		Know		Use	
	Observations	% total	Observations	% total	Observations	% Know
SMEs	610	52.1	560	47.9	102	18.2
Large firms	87	16.2	451	83.8	191	42.4
Total	697	40.8	1,011	59.2	293	29.0
Firms that conduct R&D						
	Observations	% total	Observations	% total	Observations	% Know
SMEs	62	25.6	180	74.4	82	45.6
Large firms	41	10.3	357	89.7	177	49.6
Total	103	16.1	537	83.9	259	48.2

Source: ESEE and own elaboration

⁴ See Fariñas and Jaumandreu (1999) and www.funep.es for the survey's design and elaboration.

⁵ We sort by sizes because of the special characteristics of the sample. Large firms (more than 200 employees) were fully selected, while small ones (>10 and ≤ 200 employees) were randomly selected.

⁶ Sectors are classified into high-medium and medium-low technological intensity according to www.ine.es/daco/daco42/daco4217/lstscetnae.doc.

TABLE 2.—*Firms That Know about and Use R&D Tax Incentives by Sector (2002)*

Full sample						
	Do not know		Know		Use	
	Observations	% total	Observations	% total	Observations	% Know
Medium-low tech	93	26.4	259	73.6	110	42.5
High-medium tech	604	44.5	752	55.5	183	24.3
Total	697	40.8	1,011	59.2	293	29.0
Firms that conduct R&D						
	Observations	% total	Observations	% total	Observations	% Know
Medium-low tech	25	11.7	188	88.3	106	56.4
High-medium tech	78	18.3	349	81.7	153	43.8
Total	103	16.1	537	83.9	259	48.2

Source: ESEE and own elaboration

TABLE 3.—*R&D Technological Effort of Firms That Conduct R&D (2002)*

	Do not know	Know	Use
SMEs	1.5	2.3	3.0
Large firms	0.8	1.6	2.2
Total	1.2	1.9	2.5

Source: ESEE and own elaboration

4. Methods

We examine the average treatment effects of R&D tax incentives on firms' R&D technological effort using two different approaches: matching estimators and Heckman's two-step selection model.

The main advantage of the first approach is that we need to assume neither functional form nor distribution of the measurement error term. Other advantage is that the matching procedure considers only treated and non treated units in the *common support*, by dropping all the controls whose variable's value is higher or smaller than that of the treated. Finally, this method reduces the number of non-treated to a sub-sample (the *selected controls*) with characteristics more homogeneous to those of treated units. Nonetheless, this method also has some drawbacks: it allows us only to verify heterogeneity between the treated and not-treated firms (Czarnitzki *et ál.*, 2004); it assumes that selection depends on observable factors only, which in turn means that the researcher has to observe every relevant factor that could affect the probability of receiving the treatment

(Hussinger, 2006); and it relies strongly on the existence of common support (Heckman *et al.*, 1997 and 1998).

Because the matching method allows only for a binary treatment and it is only reliable just in the case of *selection on observables*⁷, we compare the results using a second method. Selection models take into account selection on either observable or unobservable variables. A characteristic of these models is the additive separability of the unobserved components. A failure of common support is not as big a problem as for the matching approach, although this method has the need of requiring some specific distributional hypothesis than other models do not need and the cost of a formal functional form.

In the empirical analysis, we first consider the full sample of firms that report knowing about the tax incentives (which includes firms not conducting R&D), and second, the subsample of only firms that perform R&D. We work with the full sample because we take into account firms' R&D status in order to evaluate the inducement effect of tax incentives on firms that do not conduct R&D (Aerts and Czarnitzki, 2005; González and Pazó, 2008). If we judged only those that did, we could underestimate the treatment effect because we could not test whether or not a firm would have performed innovation activities without the incentives.

To realize the economic and statistical meaning of both approaches, we present a concise overview about them.

4.1. THE MATCHING APPROACH

The matching approach is used when the data exhibit a selection bias. If R&D tax incentives were randomly distributed among firms, then we could simply compare the average R&D technological effort of the user and non-user firms in order to obtain the treatment effect.⁸ But fiscal incentives for R&D are not granted randomly; they depend on the government's design. Moreover, firms at least are partly self-select into the treatment and their decisions whether or not to take advantage of these incentives may be related to the expected benefits. The matching method attempts to emulate an experiment where tax incentives are distributed randomly by comparing the outcomes of pairs of firms that are very similar in several characteristics. We can then compare the average R&D technological effort of the group of user firms with the group of matched similar non-user firms.

⁷ We have *selection on observables* when only characteristics observable to the researcher are correlated with the treatment status variable so that, controlling opportunely for them, treatment effects can be estimated consistently. Although this limit, if the researcher has a wide set of observed variables (as in our case) the problem of selection on unobservables should be attenuated.

⁸ Under the weaker assumption of mean independence, which supposes that firms' participation in the fiscal incentives system is random, we can estimate the output by simply averaging the treatment and non-treatment firms (Wooldridge, 2001).

In the matching method, we evaluate the effect of tax incentives on R&D technological effort (Y_1) for the firms that use them (treated group, $A=1$), and the potential outcome (Y_0) for the non-user firms (untreated group, $A=0$). The fundamental problem of causal inference is that it is impossible to observe the individual treatment effect. But, under some assumptions, we can estimate the average treatment effect for the sample. We calculate the average treatment effect on the treated as follows:

$$ATT = E(Y_1 - Y_0 | A = 1) = E(Y_1 | A = 1) - E(Y_0 | A = 1) \quad [1]$$

And we estimate the average treatment effect on the untreated as follows:

$$ATC = E(Y_1 - Y_0 | A = 0) = E(Y_1 | A = 0) - E(Y_0 | A = 0) \quad [2]$$

To carry out the evaluation, we need to construct the counterfactual $E(\hat{Y}_0 | A = 1)$ in equation [1], that is, the R&D technological effort users would have made, on average, had they not taken advantage of the tax incentives, selecting from the non-user firms a control group in which the distribution of observed variables is as similar as possible to the distribution in the treated group. We also need to construct the counterfactual $E(\hat{Y}_1 | A = 0)$ in equation [2], that is, the R&D technological effort non-user firms would have made, on average, had they taken advantage of the tax incentives, selecting from the user firms the control group with the most similar distribution of observed variables.

To construct the counterfactual, we assume that all relevant differences between user and non-user firms are captured by their observable characteristics, X . Then we select from the non-user (user) firms a control group in which the distribution of observed variables is as similar as possible to the distribution in the user (non-user) group. We use the propensity score (PS) proposed by Rosenbaum and Rubin (1983) instead of the vector of the observable characteristics X as the matching argument. They define the propensity score as the conditional probability of using tax incentives [$p(x) = \Pr\{A = 1 | X = x\}$]. They show that if the exposure to treatment is random within cells defined by X , it is also random within cells defined by the values of a synthetic measure $p(x)$.

After that, we pair each user firm with some group of “equivalent” non-user firms, and associate the R&D technological effort of the treated firms i , y_i with the weighted outcomes of their “neighbours” j in the untreated comparison group ($A=0$) to obtain [1]:

$$\hat{y}_i = \sum_{j \in A=0} w_{ij} y_j \quad [3]$$

And we pair each non-user firm with the group of “more similar” user firms, and associate the R&D technological effort of the untreated firms j , y_j with the weighted outcomes of their “neighbours” i in the treated comparison group ($A=1$) to obtain [2]:

w_{ij} and $w_{ji} \in [0,1]$ with $\sum_{j \in A=0} w_{ij} = \sum_{i \in A=1} w_{ji} = 1$ with are the weighted controls in forming a comparison with the other groups, untreated and treated, respectively.

We estimate two types of matching estimators to obtain the treatment effect for the treated (user) firms and the untreated (non-user) firms.⁹ First, we find the outcome of the most similar control unit, that is, *nearest neighbour matching*¹⁰

$$j : |p_i - p_j| = \min_{k \in \{A=0\}} \{|p_i - p_k|\} \quad w_{ik} = 1(K = j) \quad [5]$$

to estimate the treatment effect on the treated firm i , and

$$i : |p_j - p_i| = \min_{k \in \{A=1\}} \{|p_j - p_k|\} \quad w_{jk} = 1(K = i) \quad [6]$$

to estimate the treatment effect on the untreated firm j .

Second, we use a weighted average of the outcomes of the untreated (treated) firms where the weight given to untreated j (treated i) firms is in proportion to the nearness of the observable characteristics of i (j) and j (i) (this is Kernel-based matching)¹¹.

$$w_{ij} = \frac{K\left(\frac{p_i - p_j}{h}\right)}{\sum_{j \in A=0} K\left(\frac{p_i - p_j}{h}\right)} \quad [7]$$

to estimate the treatment effect for the treated firms i , and

⁹ See Heckman, Ichimura and Todd (1997, 1998), Heckman, Ichimura, Smith and Todd (1998) and Becker and Ichino (2002) for a discussion of different types of matching estimators.

¹⁰ Nearest neighbour matching has been used in Czarnitzki (2001), Czarnitzki and Fier (2002), Almus and Czarnitzki (2003), Kaiser (2004), Lööf and Hesmati (2005), Czarnitzki and Ebersberger (2006), González and Pazó (2008), Herrera and Heijs (2006) and Aerts and Czarnitzki (2005).

¹¹ Kernel-based matching has been used by Duguet (2004) and Kaiser (2004).

$$w_{ij} = \frac{K\left(\frac{p_j - p_i}{h}\right)}{\sum_{i \in A=1} K\left(\frac{p_j - p_i}{h}\right)} \quad [8]$$

to estimate the treatment effect for the untreated firms j .

Finally, we consider the average difference between the actual and the estimated outcome (R&D technological effort) to be caused by the tax policy. To derive equations [1] and [2] we need two assumptions given the propensity score: first, the *balancing property*, that is, the balancing of pre-treatment variables given the propensity score

$$A \perp X \mid p(X) \quad [9]$$

where A shows the treatment status (using or not using the tax incentives); and second, the *unconfoundedness assumption* given the propensity score

$$Y_1, Y_0 \perp A \mid p(x). \quad [10]$$

If the first assumption is satisfied, observations with the same propensity score must have the same distribution of observable characteristics independently of their treatment status (the second assumption). That is, for a given propensity score, exposure to treatment is random; therefore, user and non-user firms should be on average observationally identical.

We estimate the propensity score according to Becker and Ichino (2002)¹², considering the normal cumulative distribution

$$p(x) = \Pr\{A = 1 \mid X = x\} = F(h(X = x))$$

where $h(X)$ is a function of covariates with linear and higher order terms. We restrict the sample to common support; that is, we delete all the observations on user firms with probabilities larger than the maximum and smaller than the mini-

¹² The program *pscore.ado* proposed by Becker and Ichino (2002) estimates the propensity score and tests the balancing assumption.

mum in the potential control group. Also, for both samples (the full sample and those firms that perform R&D and know about the tax incentives), we estimate separate propensity scores for larger firms (more than 200 employees) and for SMEs, to evaluate the effects on the outcome (R&D technological effort) by size. So we estimate in total six propensity scores, and then we calculate the average treatment effects on the treated firms using nearest neighbour and Kernel-based matching as well as the average treatment effects on the non-treated firms using nearest neighbour matching¹³.

To estimate the propensity score we use a similar specification from Corchuelo and Martínez-Ros (2008), in which the dependent variable is a dummy with value 1 if the firm takes advantage of the tax incentives and 0 otherwise. The vector of the covariates includes variables classified in three groups. The first group considers binary variables for firms' characteristics: size (*D10-20*, *D21-50*, *D51-100*, *D101-200*, *D201-500*, *D>500*, with value 1 if the firm is in the given interval of number of employees), to take into account management abilities (Czarnitzki and Fier, 2002; González and Pazó, 2008); sector (*DM-HS*, with value 1 if the firm is in a high-medium tech sector), to capture technological opportunities; and a financial variable (*financial stability*, with value 1 if the firm's own equity over total debt is >0.5 and 0 otherwise), to account for differences in the way that firms finance their investment activities (Harhoff, 1998; Hall, 2002 and 2005).

The second group includes binary variables linked with firms' innovation activities: *stable R&D firm*, with value 1 if the firm has performed R&D activities in the period 1998-2000, and 0 if it has not performed R&D activities or has done so in only some of those years, to show the experience of the firms (Busom 2000; Blanes and Busom 2004); and *receive subsidies in 2002* to take into account that firms that participate in this type of program are also aware of other kinds of public support and have perfectly identified the R&D expenditures that qualify for the deductions.

Finally, we have considered some variables that assume that firms evaluate costs and benefits before choosing to take advantage of the tax incentives. The benefits are approximated by the variable *B-index*¹⁴; it represents the firm's minimum profit expected to obtain from investing in R&D, allowing for tax incentives¹⁵. The costs are approximated by two binary variables that reveal the obstacles perceived by managers to using tax incentives: *lack of R&D personnel*, with value 1 if the firm did not dedicate any personnel to R&D in 2002; and *lack of importance of quality improvements in production*, with value 1 if the firm does not perform quality controls, quality is not the principal reason for price varia-

¹³ We employ the method described in Becker and Ichino (2002) and Abadie and Imbens (2004).

¹⁴ The B-index is a synthetic index designed by McFetridge and Warda (1983) that allows comparing the generosity of different systems of fiscal incentives for R&D. It is also the fiscal component of the cost of capital of a marginal project of R&D.

¹⁵ The variable B-index has been elaborated in a similar way of Corchuelo and Martínez-Ros (2008).

tion, the firm does not perform market and marketing research to commercialise its product, and the firm has no design activities, and 0 otherwise.

We test¹⁶ that the variables *B-index* and *receive subsidies in 2002* are endogenous, and we use instrumental variables to take the endogeneity into account. Therefore, we include in the regressions the variable *B-index* for the previous year, and we use the estimated probability of the variable *receive subsidies in 2002* (a binary variable with value 1 if the firm obtained an R&D subsidy in 2002) with a set of covariates¹⁷.

Finally, the predicted probabilities (the so-called propensity scores) obtained from the probit models serve as argument in the matching procedure.

4.2. THE SELECTION MODEL APPROACH

We compare the performance of the previous approach with a sample selection approach (as in the Heckman two-stage selection model, 1979), which is a powerful method to deal with the selection on observables. We use this methodology implementing instrumental variables approach as well, in order to solve the problem of *selection on unobservables*. In this case, we need to know a full set of exogenous variables (the instruments) correlated with the treatment variable and (*Y*), at the same time, uncorrelated with the outcome (*R&D technological effort*), in order to build a 2SLS estimation of the evaluation equation.

The model is composed of two correlated equations, one for the outcome (*R&D technological effort* as a function of a set of variables) and one for the selection equation (the probability of using the tax incentives as a function of the previous set of variables), that takes the following form for each *j-esima* firm:

$$\left\{ \begin{array}{l} \ln(esf)_t = \gamma z_j + \alpha \ln(b - index_{t-1}) + u_j \quad u_j \rightarrow N(0, \sigma_u^2) \\ use^* = \beta x_j + \varepsilon_j \quad \varepsilon_j \rightarrow N(0, \sigma_\varepsilon^2) \\ use = \begin{cases} 1 & \text{if } use^* > 0 \\ 0 & \text{otherwise} \end{cases} \\ Cov(u_j, \varepsilon_j) = \rho \neq 0 \end{array} \right.$$

¹⁶ We use the Blundell and Smith test for probit models to check for the potential endogeneity of these variables. Under the alternative hypothesis, we regress the suspected endogenous variables on a linear projection of a set of instruments. We then include the residuals obtained in the model and, under the null hypothesis; exogeneity can be rejected, so we test the endogeneity of the variables.

¹⁷ See Corchuelo and Martínez-Ros (2008) for details about the model used.

where z and x are covariates and u and ϵ are unobservable components (error terms) with zero means but supposed to be correlated. The sign of ρ shows whether non-observables in the participation and non-observable in the outcome are positively or negatively correlated and that, to make our results the most comparable we hold $z=x$ =matching covariates (to reduce any arbitrariness in choosing different sets of x and z in the two equations). Nevertheless, we introduce another variables in the outcome equation in order to analyse the elasticity of R&D technological effort on the benefits (or economic-financial profitability) that a firm expect to get by investing in R&D. These benefits are approximated by the variable *B-index* which represents the minimum profit that one firm expects to obtain from investing in R&D, allowing for tax incentives.

5. Empirical results

5.1. PROBABILITY OF USING TAX INCENTIVES

First, we have estimated the predicted probability of using the tax incentives for the full sample (Table 4) and the sample of firms performing R&D (Table 5). In all the specifications, we have verified that the balancing property is satisfied¹⁸.

The participation probability increases with firm size when the full sample is considered, but not when we restrict the sample to those firms performing R&D. Technological opportunities are not significant in either the full sample or the subsample of innovative firms. One important factor in explaining the probability of using tax incentives is to obtain R&D subsidies. Firms that apply for R&D subsidies have perfectly identified the qualified expenditures and can support the administrative or other costs of applying for them, so, as expected, this factor influences the use of the tax incentives system. This result shows complementarities among the public instruments to promote innovation. Financial stability also increases the probability of participation, but it is significant only to SMEs showing the importance of not having financial constraints to investing in R&D. In contrast, although being a stable R&D performer increases, in general, the probability of participation, it is significant only in large firms, which shows that firms with higher capacity to innovate are also more likely to use tax incentives. Finally, results report that only large firms are aware of the benefits from the tax incentives, while SMEs emphasize on the obstacles to using them.

In sum, those Spanish firms that have higher capacity for innovation (Czarnitzki *et al.* 2004), participate in other types of public financing of innovation

¹⁸ Upon request, the authors will provide the number of blocks that ensures that the mean propensity score is not different for treated firms and control firms in each block (guaranteeing that the balancing property is satisfied) and the number of treated firms and control firms for each block considering the common support (the table that shows the inferior block of the propensity score).

activities and have a stable financing position are more likely to take advantage of tax incentives for R&D.

5.2. AVERAGE TREATMENT EFFECTS

Second, we present the average treatment effect on treated firms and the average treatment effect on untreated firms using different types of matching estimators. We have tested the full sample of firms knowing tax incentives (Table 6) and the subsample of firms that know about the tax incentives and perform R&D (Table 7), distinguishing by size. The first two rows show nearest neighbour matching (random draw version) and the Kernel matching proposed by Becker and Ichino (2002), and the rest of the rows show the sample average effect of treatment on treated firms (SATT) and the sample average effect of treatment on controls (SATC) using the nearest neighbour estimator proposed by Abadie and Imbens (2004). Effect tests are usually carried out by means of simple t-statistics. However, the ordinary t-value is biased upwards because it does not capture that the mean of the outcome variable of the control group results not from a random sampling but from an estimation (via the propensity score and the matching procedure). To remove this bias, we apply a bootstrapping method that simulates the distribution of the mean outcome of the control group by repeated sampling. This empirical distribution can calculate a standard error and, thus, a t-statistic that is not biased. On the other hand, Smith (2000) considers that the nearest neighbour estimator can operate with more than one “nearest neighbour,” and *with* or *without* replacement, that is, the untreated firm can act as counterfactual of more than one treated firm. This way, we can reduce the variance of the estimator, since the observations with insufficient “nearest neighbours” are ignored and the group of possible matching firms is larger and this can reduce biases (Abadie and Imbens 2006).

In Table 6 we observe that only the SATT procedure presents a positive and significant effect of tax incentives on R&D technological effort, either for the full sample or for any group of firms by size. Comparatively, the effect is higher in SMEs. However, the effect that tax incentives would have on non-user firms is not significant.

When only firms that conduct R&D are considered (Table 7), the positive and significant effect of tax incentives exists only in large firms. This result supports the idea that only large firms are aware of the benefits of the incentives and that SMEs, in contrast, perceive some obstacles that impede their access to these incentives. The hypothetical effect that tax incentives would have on non-users is positive and significant in the full sample of innovative firms: R&D technological effort could be increased if more firms took advantage of tax incentives.

Finally, we present the results of the second step of Heckman’s in Tables 8 and 9. The dependent variable in the regressions is the log of R&D technological effort, and we have used the log of *B-index* as an independent variable to obtain the elasticity of R&D effort due to the tax benefits. Results are similar to those from the matching procedure. Only the coefficients of the technological opportunities and *B-index* variables are significant and have the correct sign. R&D tech-

nological effort increases in high-tech firms and when the B-index is minor, but the former effect is significant only in large firms. Parametric estimation shows that only large firms and firms in high-tech industries seem to benefit from the tax incentives, and for those firms the tax policy is effective. As well, the Mill's ratio is not significant in SMEs, which shows that there is not a selection problem and that tax incentives are randomly distributed in firms of this size.

6. Conclusions

In general, private R&D innovation is a risky strategy overall when firm face to higher levels of uncertainty in terms of appropriability. Under such conditions, firms tend to reduce their efforts investing under the social optimum and the governments use fiscal incentives for R&D to encourage innovation and solve this market failure. Spain's tax policy is considered the most favourable among the OECD countries, but data report that it is little used by firms. We have used two different types of analysis, parametric (matching estimators) and nonparametric (selection model) approaches on a sample of Spanish manufacturing firms to show how the size of companies really affects the strategy to access the subsidies offered by the Spanish's Government.

Our findings are (1) Among firms that perform R&D activities, size per se does not affect the probability of using tax incentives, but they are used especially by large firms that can guarantee the viability of the projects and those that are more heavily involved in innovation activities. Moreover, tax incentives are also employed by firms that receive additional public financial support (direct subsidies) for R&D investment, in particular, small and medium-sized enterprises. Large firms consider tax incentives as a possible instrument to capture the returns from R&D investment.. We also find that the obstacles are important especially for SMEs. (2) In the estimation of the counterfactual situation, we show that on average tax policy fosters R&D technological effort, but the increase is significant only for large firms. We consider that, in the case of SMEs, a reduction of the obstacles might increase their probability of implementing innovations and therefore increasing their R&D spending. Finally, from the parametric analysis we confirm that the tax policy has a significant and great effect only in large firms.

The main conclusion is that in Spain, large firms in high-medium tech sectors are the companies that benefit most from tax incentives for innovation and those incentives do not have impact on SMEs, where the incentives are randomly distributed. Therefore, maybe the agencies responsible to stimulate the use of such incentives in order to foster the improvement of R&D private investment have to change the focus. For instance, it should be important reducing all the legal and administrative barriers that preclude SME to access and exploit the typical incentives of these policies.

References

- ABADIE, A., DRUKKER, D., LEBER HERR, J. and IMBENS G.W. (2004): Implementing matching estimators for average treatment effects in Stata. *The Stata Journal*, 4(3), 290-311.
- ABADIE, A., IMBENS, G.W. (2005): On the failure of the bootstrap for matching estimators. <http://ksghome.harvard.edu/~aabadie/bootstrap.pdf>
- (2006). Large sample properties of matching estimators for average treatment effects. *Econometrica*, 74 (1): 235-267.
- AERTS, K., CZARNITZKI, D. (2005): Using innovation survey data to evaluate R&D policy: the case of Flanders. *ZEW Discussion Paper*, n. 04-55.
- ALMUS, S., CZARNITZKI, D. (2003): The effects of public R&D subsidies on firm's innovation activities: the case of Eastern Germany. *Journal of Economic Statistics*, 21(1): 226-236.
- ALI-YRKKÖ, J. (2005a): Impact of public R&D financing on employment, *ENEPRI Working Paper* n°. 39/agosto.
- (2005b): Impact of public R&D financing on private R&D: Does financial constraints matter?, (Helsinki: ETLA, The Research Institute of the Finnish Economy, *Discussion Paper*, n. 94, 2004) , *ENEPRI Working Paper* n°. 30/ febrero.
- ARROW, K. (1962): Economic Welfare and the Allocation of Resources for Inventions. In R. Nelson (Ed.), *The Rate and Direction of Inventive Activity*. Princeton University Press.
- ASMUSSEN, R., BERRIOT, C. (1993) : Le credit d'impôt recherche, cout et effet incitatif. Ministère de l'Economie et des Finances, Paris.
- BECKER, S., ICHINO, A. (2002): Estimation for average treatment effects based on propensity scores. *The Stata Journal*, 2(4): 358-377.
- BÉRUBÉ Ch. y P. MOHNEN (2009): Are Firms that Receive R&D Subsidies More Innovative?. *Canadian Journal of Economics*, 42(1): 206-225
- BLANES, J.V., BUSOM, I. (2004): Who participates in R&D subsidy programs? The case of Spanish manufacturing firms. *Research Policy*, 33: 1459-1476.
- BLOOM, N., GRIFFITH, R., VAN REENEN, J. (2002): Do R&D credits work? Evidence from an international panel of countries 1979-1997. *Journal of Public Economics*, 85: 1-31.
- BUSOM, I. (2000): An empirical evaluation of the effects of R&D subsidies. *Economics of Innovation and New Technologies*, 9: 111-148.
- COLLINS, E. (1983): *An early assessment of three R&D tax incentives provided by the Economic Recovery Tax Act of 1981*, Vol. PRA Report 83-7, Washington, DC: National Science Foundation, 1983.
- CORCHUELO, M. B. (2006): Incentivos fiscales en I+D y decisiones de innovación. *Revista de Economía Aplicada*, XIV (40): 5-34.
- and MARTÍNEZ-ROS, E. (2008): Aplicación de los incentivos fiscales a la inversión en I+D en las empresas españolas, *Hacienda Pública Española/Revista de Economía Aplicada*, 187-(4/2008): 9-39.
- CZARNITZKI, D. (2001): Die auswirkungen der forschungs-und technologiepolitik auf die innovationsaktivitäten ostdeutscher unternehmen ("The impact of R&D policies on the innovation activities of Eastern German enterprises"), *Journal of Applied Social Science Studies*, 121(4): 539-560.
- CZARNITZKI, D., FIER A. (2002): Do innovation subsidies crowded out private investment: evidence from the German service sector. *Applied Economics Quarterly*, 48(1): 1-25.
- CZARNITZKI, D., HANEL, P., ROSA, J. M. (2004): Evaluating the impact of R&D tax credit on innovation: A microeconomic study on Canadian firms. *ZEW Discussion Paper* N°. 04-77.

- CZARNITZKI, D. y B. EBERSBERGER (2006): Distributional effects of public R&D funding, *mimeo*.
- DAVID, P. A., HALL, B.H., TOOLE A. A. (2000): Is public R&D a complement or substitute for private R&D? A review of the econometric evidence. *Research Policy*, 29: 497-529.
- DEHEJIA, R., WAHBA, S. (1999): Casual effects in nonexperimental studies: reevaluating the evaluation of training programmes. *Journal of the American Statistical Association*.
- (2002): Propensity Score-Matching methods for nonexperimental casual studies. *Review of Economics and Statistics*, 84(1): 151-161.
- DEPARTMENT OF FINANCE CANADA AND REVENUE CANADA (1998): *Evaluation report: The Federal system of income tax incentives for scientific research and experimental development*. Ottawa.
- DUCH, N., MONTOLIA, D., MEDIAYILLA, M. (2007): Evaluation of public subsidies oriented to firm's performance: a quasi-experimental approach. *Instituto D'Economia de Barcelona, Working Paper* 2007/3.
- DUGUET, E. (2004): Are R&D subsidies a substitute or a complement to privately funded R&D? Evidence from France using propensity score methods for non-experimental data. *Review d'Economie Politique*, 114 (2): 263-292.
- EUROPEAN COMMISSION (2003). *Raising EU R&D Intensity. Improving the Effectiveness of Public Support Mechanism for Private Sector Research and Development: Fiscal Measures*. EUR 20714, DG for Research Knowledge Based Society and Economy Strategy and Policy, Investment in Research, Luxemburgo.
- FALK, M. (2004): *European Productivity, Innovation and Public Sector R&D: Evidence from OECD country and industry-level data*, background report for the European Competitiveness Report 2004.
- FARIÑAS, J. C., JAUMANDREU, J. (1999): Diez años de Encuesta sobre Estrategias Empresariales (ESEE). *Economía Industrial*, 329: 29-42.
- FIER, A. (2002): Technology policy and business innovation activities in the case of Germany. In Gretzmacher, N., G. Hutschenreiter y W. Polt (eds.), *Proceedings of the tip workshop: Changing Strategies for Business R&D and their implications for Science and Technology* (pp.54-67).
- GELABERT, L., FOSFURI, A. TRIBÓ, J. (2009): Does the effect of public support for R&D depend on the degree of appropriability?, *Journal of Industrial Economics*, 57(4): 736-767.
- GONZÁLEZ, X., JAUMANDREU, J. y C. Pazó (2005): Barriers to innovation and subsidy effectiveness, *Rand of Journal Economics*, 36(4): 930-950.
- GONZÁLEZ, X., PAZÓ, C. (2008): Do public subsidies stimulate private R&D spending?. *Research Policy*, 37: 371-389.
- GÖRG, H. y E. STROLB (2005): The effects of R&D subsidies on private R&D, *Research Paper Series (Globalisation, Productivity and Technology)* n° 38, The University of Nottingham. *Economica*, 74: 215-234.
- GUELLEC, D., VAN POTTLESBERGHE, B. (2003): The impact of public R&D expenditure on business R&D. *Economics of Innovation and New Technologies*, 12(3): 225-244.
- HALL, B. (1993). R&D tax policy during the 1980s: success or failure. *Tax Policy and the Economy*, 7: 1-35.
- (2002). The financing of research and development. *Oxford Review of Economic Policy*, 18(1): 35-51.
- (2005): The financing of innovation, en Shane, S. (ed.), *Blackwell Hand book of Technology and Innovation Management*, Oxford: Blackwell Publishers, Ltd., 2005.
- and VAN REENEN, J. (2000): How effective are fiscal incentives for R&D? A review of the evidence. *Research Policy*, 29: 449-469.

- HANEL, P. (2003): Impact of Government support programs on innovation by Canadian manufacturing firms. *Note de Recherche* 2003-09, CIRST: Centre Interuniversitaire de recherche sur la science et la technologie (CIRST).
- HARHOFF, D. (1998): Are the financing constraints for R&D investment in German manufacturing firms?. *Annales D'Économie et de Statistique*, 49/50: 421-456.
- HECKMAN, J., ICHIMURA, H., TODD, P. (1997): Matching as an econometric evaluation estimator: evidence from evaluating a job training program. *Review of Economic Studies*, 64 (5): 605-664.
- HECKMAN, J. J. (1979): Sample selection bias as a specification error, *Econometrica*, 47: 153-161.
- HECKMAN, J., ICHIMURA, H., TODD, P. (1998): Characterizing selection bias using experimental data. *Econometrica*, 66 (5): 1017-1098.
- HECKMAN, J., LALONDE, R., SMITH, J. (1999): The economics and econometrics of active labour market programme. In Ashenfelter, O. and Card, D. (eds.), *The Handbook of Labor Economics*, Vol. III.
- HEIJIS, J., M. BUESA, L. HERRERA and P. VALADEZ (2006): Evaluación de los incentivos fiscales a la I+D+i en España basado en el propensity score matching, *Documento de Investigación*, nº. 3, julio, Instituto de Estudios Fiscales.
- HERRERA, L. y J. HEIJIS (2006): Difusión y adicionalidad de las ayudas públicas a la innovación, *Revista de Economía Aplicada*, 14(41):. 45-65.
- HOTZ, V. J., IMBENS, G.W., KLERMAN, J. A. (2006): Evaluating the differential effects of alternative welfare-to-work training components: a re-analysis of the Californian GAIN program. *NBER Working Paper* nº. 11939.
- HUSSINGER, K. (2006): R&D and subsidies at the firm level: An application of parametric and semi-parametric two-step selection models, *Catholic University Leuven and Centre for European Economic Research (ZEW)*.
- KAISER, U. (2004): Private R&D and public subsidies: microeconomic evidence from Denmark. *Center for Economic and Business Research*, DP 2004-19.
- LACH, S. (2002): Do R&D subsidies stimulate or displace private R&D? Evidence from Israel, *The Journal of Industrial Economics*, 4, December, pp. 369-390.
- LATTIMORE, R. (1997): Research and Development Fiscal Incentives in Australia: Impacts and Policy Lessons. Industry Commission.
- LECHNER, M. (2002): Some practical issues in the evaluation of heterogeneous labour market programmes by matching methods. *Journal of the Royal Statistical Society: Serie A (Statistics in Society)*, 165, 59.
- LÖÖF, H., HESMATI, A. (2005): Additionality or crowding out? On the effectiveness of R&D subsidies. *The Royal Institute of Technology. Centre of Excellence for Studies in Science and Innovation*.
- MANSFIELD, E. (1986): The R&D tax credit and other technology policy issues. *American Economic Review*, 76 (2): 190-194.
- MARRA, M. A. (2004): Incentivos fiscales, inversión en actividades de I+D y estructura de costes. Un análisis por tamaño para una muestra representativa de empresas manufactureras españolas 1991-1999. *Hacienda Pública Española/ Revista de Economía Pública*, 170-(3/2004): 9-35.
- MC FETRIDGE, D. G., WARDA, J.P. (1983): Canadian R&D Tax Incentives: Their Adequacy and Impact. *Canadian Tax Paper*, 70. Toronto: Canadian Tax foundation.
- NELSON, R. (1959): The simple economics of basic scientific research. *Journal of Political Economy*, 76: 297-306.
- PARISI, M. L., SEMBENELLI, A. (2003): Is private R&D spending sensitive to its price? Empirical evidence on panel data for Italy. *Empirica*, 4.
- Romero, D., Sanz, J.F. (2007): Eficacia de los incentivos fiscales a la inversión en I+D en los años noventa, *Hacienda Pública Española. Revista de Economía Pública*, 183 (4): 9-32.

- ROSEBAUM, P. R., RUBIN, D. B. (1983): The central role of the propensity score in observational studies for causal effects", *Biometrika*, 70 (1): 41-55.
- RUBIN, D.B. (1977): Assignment to treatment group on the basis of covariate, *Journal of Educational Statistics* 2: 1-26.
- SIANESI, B. (2001): Implementing propensity score matching estimators with Stata. <http://fmwww.bc.edu/RePec/usug2001/psmatch.pdf>
- SMITH, J. (2000): A critical survey of empirical methods for evaluating active labor market policies, *Zeitschrift für Volkswirtschaft und Statistik*, 136: 1-22.
- SUETENS, S. (2002): R&D subsidies and production effects of R&D personnel: evidence from the Flemish region, *CEST Discussion Paper* 2002/03, Antwerp.
- TOIVANEN, O. y NIININEN, P. (2000): Investment, R&D, subsidies and credit constraints, *Working Paper W-264*, Helsinki School of Economics and Business Administration, Helsinki.
- VAN POTTLESBERGHE, B., MEGALLY, E., NYSTEN, S. (2003): Evaluation of current fiscal incentives for business R&D in Belgium. *Working Paper CEB 011*. Solvay Business School, Université Libre de Bruxelles.
- WALLSTEN, S. (2000): The effect of government-industry R&D programs on private R&D: the case of small business innovation research policy, *Rand Journal of Economics*, 1: 82-100.
- WARDA, J. (2001): Measuring the value of R&D tax treatment in OECD countries. In *SIT Review*, 27, 185-211.
- WARDA, J. (2002): A 2001-2002 update of R&D tax treatment in OECD countries. Paper prepared for the OECD Directorate for Science, Technology and Industry.
- WOOLDRIDGE, J. M. (2001): *Econometric Analysis of Cross Section and Panel Data*. The MIT Press, Cambridge, Ma.

TABLE 4.—*Probability of Using Tax Incentives, among Firms That Know about Them*
[Dependent variable: *claim R&D deductions*]

	Total	SMEs	Large
Const.	-1.80 (-5.4)	-0.9 (-3.0)	-0.9 (-3.8)
D21-50*	0.71 (2.4)		
D51-100*	0.69 (2.1)		
D101-200*	1.04 (3.4)		
D201-500 *	0.94 (3.1)		
D> 500 *	0.72 (2.4)		
DM-HS *	0.11 (1.0)	0.08 (0.4)	0.07 (0.5)
Financial stability*	0.30 (3.0)	0.43 (2.9)	0.21 (1.5)
Stable R&D firm*	0.62 (3.9)	0.45 (1.8)	0.66 (3.2)
$\hat{P}_{\text{(receive subsidies in 2002)}}$	0.69 (3.0)	1.73 (2.8)	0.51 (2.0)
$\hat{B} - index_{t-1}$	0.65 (2.3)	0.57 (1.0)	0.72 (2.2)
Lack of R&D personnel*	-0.49 (-3.1)	-0.80 (-3.2)	-0.38 (-1.8)
Importance of quality improvements in production *	-0.40 (-3.1)	-0.48 (-2.5)	-0.34 (-1.8)
Number of obs.	914	518	396
Log-likelihood	-409.1	-181.9	-231.33
Region of common support	[0.018, 0.92]	[0.023,0.96]	[0.111,0.83]

*Note: $\frac{dy}{dx}$ is for discrete change of dummy variable from 0 to 1; in parenthesis, *t* statistics robust to heteroscedaticity.

TABLE 5.—*Probability of Using Tax Incentives, among Firms That Conduct R&D and Know about the Incentives*
 [Dependent variable: *claim R&D deductions*]

	Total	SMEs	Large
Const.	-1.30 (-2.7)	-0.80 (-2.3)	-1.10 (-3.9)
D21-50*	0.72 (1.5)		
D51-100*	0.88 (1.7)		
D101-200*	0.91 (1.8)		
D201-500*	0.81 (1.7)		
D> 500*	0.56 (1.1)		
DM-HS*	0.23 (1.7)	0.27 (1.1)	0.09 (0.6)
Financial stability*	0.26 (2.2)	0.47 (2.2)	0.14 (0.9)
Stable R&D firm*	-	0.13 (0.5)	0.80 (3.4)
$\hat{P}_{\text{(receive subsidies in 2002)}}$	0.81 (3.5)	1.55 (2.4)	0.52 (2.0)
$\hat{B} - index_{t-1}$	0.79 (2.7)	0.73 (1.2)	0.75 (2.2)
Lack of R&D personnel*	-0.42 (-2.1)	-0.05 (-0.2)	-0.63 (-2.3)
Importance of quality improvements in production*	-0.31 (-1.9)	-0.46 (-1.8)	-0.19 (-0.8)
Number of obs.	487	171	316
Log-likelihood	-307.2	-101.9	-194.6
Region of common support	[0.120, 0.92]	[0.169, 0.97]	[0.062, 0.83]

*Note: $\frac{dy}{dx}$ is for discrete change of dummy variable from 0 to 1; in parenthesis, t statistics robust to heteroscedasticity.

TABLE 6.—*Matching Estimators: Firms that Know about the Tax Incentives*

	Total	SMEs	Large
ATT (NNM)	0.35 (0.9) ⁽¹⁾	0.56 (0.9) ⁽¹⁾	0.84 (2.3)⁽¹⁾
N. Treated	283	101	182
N. Controls	149	61	84
ATT (Kernel)	0.76 (3.2)⁽¹⁾	0.53 (1.2) ⁽¹⁾	0.74 (2.6)
N. Treated	283	101	182
N. Controls	149	390	188
SATT (NNM)	0.67 (2.6)	0.87 (2.1)	0.68 (2.3)
N. Treated	283	101	182
N. Controls	615	417	198
Percent of exact matching	100	100	100
Number of matches	4	4	4
SATC (NNM)	0.23 (1.1)	0.16 (0.7)	0.33 (1.2)
N. Treated	615	417	198
N. Controls	283	101	182
Percent of exact matching	100	100	100
Number of matches	4	4	4

Note: (1) Bootstrap standard errors (n. reps: 100)

TABLE 7.—*Matching Estimators: Firms that Conduct R&D and Know about the Tax Incentives*

	Total	SMEs	Large
ATT (NNM)	0.12 (0.3) ⁽¹⁾	0.47 (0.8) ⁽¹⁾	0.68 (1.8) ⁽¹⁾
N. Treated	249	81	168
N. Controls	114	42	77
ATT (Kernel)	0.75 (2.6)⁽¹⁾	0.63 (1.1) (1)	0.79 (2.6)
N. Treated	249	81	168
N. Controls	231	85	142
SATT (NNM)	0.84 (2.6)	0.76 (1.2)	0.82 (2.3)
N. Treated	249	81	182
N. Controls	220	87	131
Percent of exact matching	100	100	100
Number of matches	4	4	4
SATC (NNM)	0.59 (2.0)	0.71 (1.3)	0.44 (1.2)
N. Treated	220	87	131
N. Controls	249	81	168
Percent of exact matching	100	100	100
Number of matches	4	4	4

Note: (1) Bootstrap standard errors (n. reps: 100)

TABLE 8.—*Heckman's Second Step (Structural Equation). Firms that Know about the Tax Incentives*
[Dependent variable: log (R&D technological effort)]

	Total	SMEs	Large
D21-50*	0.51 (0.9)		
D51-100*	-0.09 (-0.2)		
D101-200*	-0.33 (-0.6)		
D201-500*	-0.60 (-1.2)		
D> 500*	-0.72 (-1.2)		
DM-HS*	0.79 (4.2)	0.70 (2.2)	0.78 (3.0)
Financial stability*	-0.18 (-1.0)	-0.48 (-1.6)	-0.08 (-0.4)
Stable R&D firm*	0.43 (1.2)	0.51 (1.7)	0.06 (0.2)
\hat{P} (receive subsidies in 2002)	0.63 (2.0)	0.75 (1.2)	0.40 (1.0)
$\text{Ln}(\hat{B} - \text{index}_{t-1})$	-0.43 (-2.1)	-0.18 (-0.6)	-0.78 (-2.4)
Mill's ratio	-0.97 (-3.3)	-0.60 (-2.5)	-1.71 (-5.2)
Number of obs.	243	81	162
Wald chi (free degrees)	134.8(19)	56.0(9)	63.9(9)

Note: * $\frac{dy}{dx}$ is for discrete change of dummy variable from 0 to 1; in parenthesis, t statistics robust to heteroscedaticity.

TABLE 9.—*Heckman's Second Step (Structural Equation). Firms that Conduct R&D and Know about the Tax Incentives*
[Dependent variable: log (R&D technological effort)]

	Total	SMEs	Large
D21-50*	0.30 (0.6)		
D51-100*	-0.22 (-0.3)		
D101-200*	-0.41 (-0.7)		
D201-500*	-0.64 (-1.0)		
D> 500*	-0.7 (-1.0)		
DM-HS*	0.75 (3.7)	0.67 (2.0)	0.77 (2.6)
Financial stability*	-0.23 (-1.2)	-0.54 (-1.7)	-0.09 (-0.4)
Stable R&D firm*	0.71 (2.0)	0.71 (2.0)	0.25 (0.8)
\hat{P} (receive subsidies in 2002)	0.49 (1.3)	0.67 (1.0)	0.32 (0.7)
$\text{Ln}(\hat{B} - \text{index}_{t-1})$	-0.49 (-2.0)	-0.02 (-0.1)	-0.89 (-2.3)
Mill's ratio	-1.28 (-2.5)	-0.56 (-1.2)	-2.02 (-4.5)
Number of obs.	243	81	162
Wald chi (free degrees)	104.0(19)	42.2(9)	47.3(9)

Note: * $\frac{dy}{dx}$ is for discrete change of dummy variable from 0 to 1; in parenthesis, t statistics robust to heteroscedaticity.

TABLE A.1.—*Descriptive statistics*

Variables	Total		Firms That Perform R&D		Firms That Know Incentives	
	Media	s.d.	Media	s.d.	Media	s.d.
D10-20*	0.22	0.41	0.04	0.19	0.12	0.32
D21-50*	0.27	0.44	0.13	0.34	0.21	0.41
D51-100*	0.10	0.30	0.08	0.28	0.10	0.30
D101-200*	0.10	0.30	0.12	0.32	0.12	0.33
D201-500*	0.20	0.40	0.38	0.48	0.27	0.44
D > 500*	0.12	0.32	0.25	0.43	0.18	0.38
DM-HS*	0.21	0.41	0.33	0.47	0.25	0.44
Financial stability >0.5*	0.38	0.48	0.41	0.49	0.41	0.49
Stable R&D firms*	0.30	0.46	0.81	0.40	0.44	0.50
Receive R&D subsidies in 2002*	0.09	0.29	0.25	0.44	0.15	0.36
$\hat{B} - index_{t-1}$	0.41	0.14	0.45	0.21	0.42	0.17
Lack of R&D personnel *	0.68	0.47	0.14	0.35	0.53	0.50
Importance of quality improvements in production*	0.43	0.50	0.17	0.38	0.31	0.46
R&D technological effort						